

CE 530 Molecular Simulation

Lecture 10 Simple Biasing Methods

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Review

○ Monte Carlo simulation

- *Markov chain to generate elements of ensemble with proper distribution*

○ Metropolis algorithm

- *relies on microscopic reversibility $\pi_i \pi_{ij} = \pi_j \pi_{ji}$*
- *two parts to a Markov step*
 - generate trial move (underlying transition probability matrix)
 - decide to accept move or keep original state

○ Determination of acceptance probabilities

- *detailed analysis of forward and reverse moves*
- *we examined molecule displacement and volume-change trials*

Performance Measures

- How do we improve the performance of a MC simulation?
 - *characterization of performance*
 - *means to improve performance*
- Return to our consideration of a general Markov process
 - *fixed number of well defined states*
 - *fully specified transition-probability matrix*
 - *use our three-state prototype*
- Performance measures
 - *rate of convergence*
 - *variance in occupancies*

$$\Pi \equiv \begin{pmatrix} \pi_{11} & \pi_{12} & \pi_{13} \\ \pi_{21} & \pi_{22} & \pi_{23} \\ \pi_{31} & \pi_{32} & \pi_{33} \end{pmatrix}$$

Rate of Convergence 1.

○ What is the likely distribution of states after a run of finite length?

- *Is it close to the limiting distribution?*

*Probability of being in state 3
after n moves, beginning in state 1*

$$\pi_1^{(n)} = \pi_1^{(0)} \Pi^n \equiv (1 \ 0 \ 0) \begin{pmatrix} \pi_{11}^{(n)} & \pi_{12}^{(n)} & \pi_{13}^{(n)} \\ \pi_{21}^{(n)} & \pi_{22}^{(n)} & \pi_{23}^{(n)} \\ \pi_{31}^{(n)} & \pi_{32}^{(n)} & \pi_{33}^{(n)} \end{pmatrix} = \left(\pi_{11}^{(n)} \quad \pi_{12}^{(n)} \quad \boxed{\pi_{13}^{(n)}} \right)$$

- *We can apply similarity transforms to understand behavior of Π^n*
 → *eigenvector equation* $\Pi\Phi = \Phi\Lambda \Rightarrow \Pi = \Phi\Lambda\Phi^{-1}$

eigenvalue matrix: $\Lambda = \begin{pmatrix} \lambda_1 & 0 & 0 \\ 0 & \lambda_2 & 0 \\ 0 & 0 & \lambda_3 \end{pmatrix}$

eigenvector matrix $\Phi = \left(\begin{pmatrix} \phi_1 \\ \phi_2 \\ \phi_3 \end{pmatrix} \begin{pmatrix} \phi_2 \\ \phi_2 \\ \phi_3 \end{pmatrix} \begin{pmatrix} \phi_3 \\ \phi_3 \\ \phi_3 \end{pmatrix} \right)$

Rate of Convergence 1.

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$$\pi_1^{(n)} = \pi_1^{(0)} \Pi^n \equiv (1 \quad 0 \quad 0) \begin{pmatrix} \pi_{11}^{(n)} & \pi_{12}^{(n)} & \pi_{13}^{(n)} \\ \pi_{21}^{(n)} & \pi_{22}^{(n)} & \pi_{23}^{(n)} \\ \pi_{31}^{(n)} & \pi_{32}^{(n)} & \pi_{33}^{(n)} \end{pmatrix} = \left(\pi_{11}^{(n)} \quad \pi_{12}^{(n)} \quad \boxed{\pi_{13}^{(n)}} \right)$$

- *We can apply similarity transforms to understand behavior of Π^n*
 → *eigenvector equation* $\Pi\Phi = \Phi\Lambda \Rightarrow \Pi = \Phi\Lambda\Phi^{-1}$

$$\begin{aligned} \Pi^n &= (\Phi\Lambda\Phi^{-1})(\Phi\Lambda\Phi^{-1}) \dots \text{(n times)} \dots \Phi^{-1} \\ &= \Phi\Lambda(\Phi^{-1}\Phi)\Lambda(\Phi^{-1}\Phi) \dots \\ &= \Phi\Lambda^n\Phi^{-1} \end{aligned} \quad \Lambda^n = \begin{pmatrix} \lambda_1^n & 0 & 0 \\ 0 & \lambda_2^n & 0 \\ 0 & 0 & \lambda_3^n \end{pmatrix}$$

Rate of Convergence 2.

- Likely distribution after finite run

$$\pi_i^{(n)} = \pi_i^{(0)} \Pi^n = \pi_i^{(0)} \Phi \Lambda^n \Phi^{-1}$$

$$\Lambda^n = \begin{pmatrix} \lambda_1^n & 0 & 0 \\ 0 & \lambda_2^n & 0 \\ 0 & 0 & \lambda_3^n \end{pmatrix} = \begin{pmatrix} 1 & 0 & 0 \\ 0 & \lambda_2^n & 0 \\ 0 & 0 & \lambda_3^n \end{pmatrix} \xrightarrow{n \rightarrow \infty} \begin{pmatrix} 1 & 0 & 0 \\ 0 & 0 & 0 \\ 0 & 0 & 0 \end{pmatrix}$$

- Convergence rate determined by magnitude of other eigenvalues

- *very close to unity indicates slow convergence*

Occupancy Variance 1.

- Imagine repeating Markov sequence many times ($L \rightarrow \infty$), each time taking a fixed number of steps, M

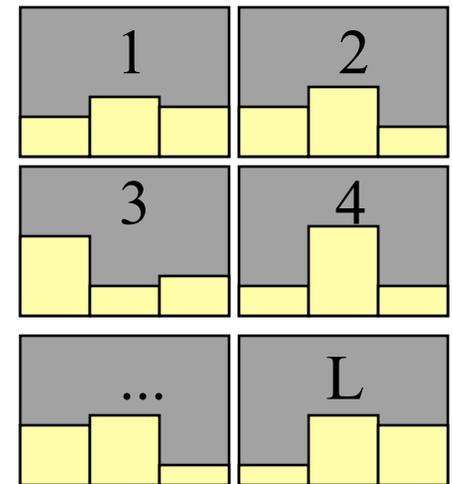
- tabulate histogram for each sequence; $p_i^{(k)} = \frac{m_i^{(k)}}{M}$
- examine variances in occupancy fraction

$$\overline{\sigma_i^2} = \sum_{k=1}^L \left(p_i^{(k)} - \pi_i \right)^2 \quad \overline{\sigma_i \sigma_j} = \sum_{k=1}^L \left(p_i^{(k)} - \pi_i \right) \left(p_j^{(k)} - \pi_j \right)$$

- through propagation of error, the occupancy (co)variances sum to give the variances in the ensemble averages; e.g. (for a 2-state system)

$$\sigma_U^2 = U_1^2 \overline{\sigma_1^2} + U_2^2 \overline{\sigma_2^2} + 2U_1 U_2 \overline{\sigma_1 \sigma_2}$$

- we would like these to be small



Occupancy Variance 2.

- A formula for the occupancy (co)variance is known

$$M \overline{\sigma_i^2} = \pi_i^2 + 2\pi_i s_{ii} - 1 \quad \text{variance}$$

$$M \overline{\sigma_i \sigma_j} = \pi_i \pi_j + \pi_i s_{ij} + \pi_j s_{ji} \quad \text{covariance}$$

$$S = (\mathbf{I} - \mathbf{\Pi} + \mathbf{\Phi})^{-1} - \mathbf{\Phi}$$

- *right-hand sides independent of M*
- *standard deviation decreases as $1/\sqrt{M}$*

Example Performance Values

Limiting distribution $\pi = (0.25 \ 0.5 \ 0.25)$

Inefficient $\Pi = \begin{pmatrix} 0.97 & 0.02 & 0.01 \\ 0.01 & 0.98 & 0.01 \\ 0.01 & 0.02 & 0.97 \end{pmatrix}$ $\lambda = \begin{pmatrix} 1 & 0.96 & 0.96 \end{pmatrix}$ $\Sigma = \begin{pmatrix} 9.2 & -6.1 & -3.1 \\ -6.1 & 12.2 & -6.1 \\ -3.1 & -6.1 & 9.2 \end{pmatrix}$

Barker $\Pi = \begin{pmatrix} 0.42 & 0.33 & 0.25 \\ 0.17 & 0.66 & 0.17 \\ 0.25 & 0.33 & 0.42 \end{pmatrix}$ $\lambda = \begin{pmatrix} 1 & 0.33 & 0.17 \end{pmatrix}$ $\Sigma = \begin{pmatrix} 0.30 & -0.25 & -0.05 \\ -0.25 & 0.50 & -0.25 \\ -0.05 & -0.25 & 0.30 \end{pmatrix}$

Most efficient $\Pi = \begin{pmatrix} 0 & 1 & 0 \\ 0.5 & 0 & 0.5 \\ 0 & 1 & 0 \end{pmatrix}$ $\lambda = \begin{pmatrix} 1 & 0 & -1 \end{pmatrix}$ $\Sigma = \begin{pmatrix} 0.125 & 0 & -0.125 \\ 0 & 0 & 0 \\ -0.125 & 0 & 0.125 \end{pmatrix}$

Metropolis $\Pi = \begin{pmatrix} 0.0 & 0.5 & 0.5 \\ 0.25 & 0.5 & 0.25 \\ 0.5 & 0.5 & 0.0 \end{pmatrix}$ $\lambda = \begin{pmatrix} 1 & 0 & -0.5 \end{pmatrix}$ $\Sigma = \begin{pmatrix} 0.10 & -0.125 & 0.02 \\ -0.125 & 0.25 & -0.125 \\ 0.02 & -0.125 & 0.10 \end{pmatrix}$

Example Performance Values

$$\Pi = \begin{pmatrix} 0 & 0.99 & 0.01 & 0 \\ 0.99 & 0 & 0 & 0.01 \\ 0 & 0.01 & 0 & 0.99 \\ 0 & 0.01 & 0.99 & 0 \end{pmatrix}$$

Lots of movement $1 \rightarrow 2; 3 \rightarrow 4$

Little movement $(1,2) \rightarrow (3,4)$

$$\pi = (0.25 \quad 0.25 \quad 0.25 \quad 0.25)$$

Limiting distribution

$$\lambda = (1 \quad 0.98 \quad -0.99 \quad -0.99)$$

Eigenvalues

$$\Sigma = \begin{pmatrix} 6.2 & 6.2 & -6.2 & -6.2 \\ 6.2 & 6.2 & -6.2 & -6.2 \\ -6.2 & -6.2 & 6.2 & 6.2 \\ -6.2 & -6.2 & 6.2 & 6.2 \end{pmatrix}$$

Covariance matrix

Heuristics to Improve Performance

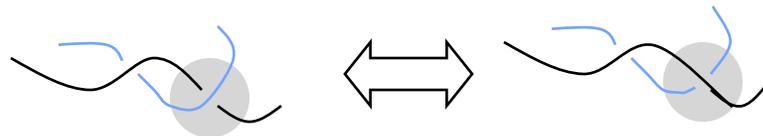
○ Keep the system moving

- *minimize diagonal elements of probability matrix*
- *avoid repeated transitions among a few states*

○ Typical physical situations where convergence is poor

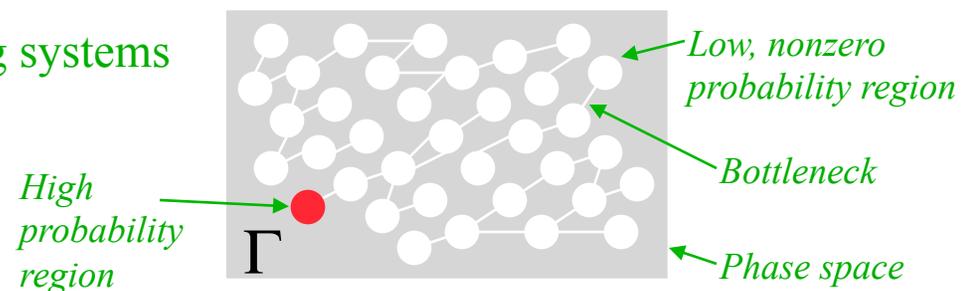
- *large number of equivalent states with poor transitions between regions of them*

→ entangled polymers



- *large number of low-probability states and a few high-probability states*

→ low-density associating systems



Biasing the Underlying Markov Process

- Detailed balance for trial/acceptance Markov process
 - $\pi_i \tau_{ij} \min(1, \chi) = \pi_j \tau_{ji} \min(1, 1/\chi)$
- Often it happens that τ_{ij} is small while χ is large (or vice-versa)
 - *even if product is of order unity, π_{ij} will be small because of min()*
- The underlying TPM can be adjusted (biased) to enhance movement among states

- *bias can be removed in reverse trial probability, or acceptance*
- *require in general*

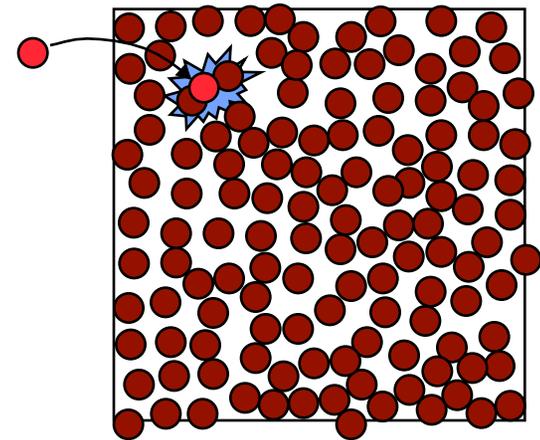
$$\chi = \frac{\pi_j \tau_{ji}}{\pi_i \tau_{ij}}$$

- *ideally, χ will be unity (all trials accepted) even for a “large” change*
 - rarely achieve this level of improvement
- *requires coordination of forward and reverse moves*

Example: Biased Insertion in GCMC

○ Grand-canonical Monte Carlo (μVT)

- *fluctuations in N require insertion/deletion trials*
- *at high density, insertions may be rarely accepted*
 - τ_{ij} is small for j a state having additional but non-overlapping molecule
- *at high chemical potential, limiting distribution strongly favors additional molecules $\pi \propto e^{\beta\mu N}$*
 - χ is large for $(N+1)$ state with no overlap
- *apply biasing to improve acceptance*
- *first look at unbiased algorithm*

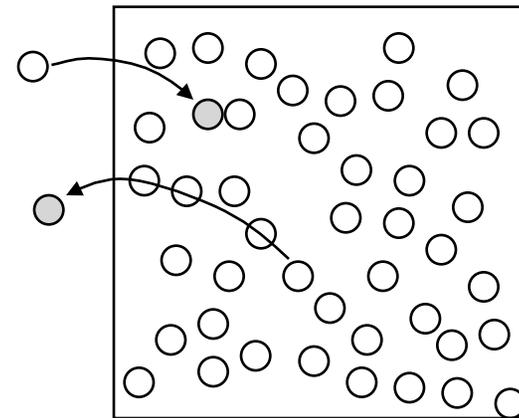


Insertion/Deletion Trial Move

1. Specification

- Gives new configuration of same volume but different number of molecules
- Choose with equal probability:
 - *insertion trial: add a molecule to a randomly selected position*
 - *deletion trial: remove a randomly selected molecule from the system*
- Limiting probability distribution
 - *grand-canonical ensemble*

$$\pi(\mathbf{r}^N) = \frac{1}{\Xi} \frac{1}{\Lambda^{dN}} e^{-\beta U(\mathbf{r}^N) + \beta \mu N} d\mathbf{r}^N$$



Insertion/Deletion Trial Move

2. Analysis of Trial Probabilities

- Detailed specification of trial moves and and probabilities

Event [reverse event]	Probability [reverse probability]	
Select insertion trial [select deletion trial]	$\frac{1}{2}$ [$\frac{1}{2}$]	<i>Forward-step trial probability</i> $\frac{1}{2} \times \frac{dr}{V} \times \min(1, \chi)$
Place molecule at \mathbf{r}_{N+1} [delete molecule N+1]	dr/V [$1/(N+1)$]	
Accept move [accept move]	$\min(1, \chi)$ [$\min(1, 1/\chi)$]	<i>Reverse-step trial probability</i> $\frac{1}{2} \times \frac{1}{N+1} \times \min(1, \frac{1}{\chi})$

χ is formulated to satisfy detailed balance

Insertion/Deletion Trial Move

3. Analysis of Detailed Balance

*Forward-step
trial
probability*

$$\frac{1}{2} \times \frac{d\mathbf{r}}{V} \times \min(1, \chi)$$

*Reverse-step
trial
probability*

$$\frac{1}{2} \times \frac{1}{N+1} \times \min(1, \frac{1}{\chi})$$

Detailed balance

$$\pi_i \pi_{ij} = \pi_j \pi_{ji}$$

*Limiting
distribution*

$$\pi(\mathbf{r}^N) = \frac{1}{\Xi} \frac{1}{\Lambda^{dN}} e^{-\beta U(\mathbf{r}^N) + \beta \mu N} d\mathbf{r}^N$$

Insertion/Deletion Trial Move

3. Analysis of Detailed Balance

Forward-step trial probability $\frac{1}{2} \times \frac{d\mathbf{r}}{V} \times \min(1, \chi)$

Reverse-step trial probability $\frac{1}{2} \times \frac{1}{N+1} \times \min(1, \frac{1}{\chi})$

Detailed balance

$$\pi_i \pi_{ij} = \pi_j \pi_{ji}$$

$$\frac{e^{-\beta U^{old} + \beta \mu N} d\mathbf{r}^N}{\Xi \Lambda^{dN}} \left[\frac{1}{2} \times \frac{d\mathbf{r}}{V} \times \min(1, \chi) \right] = \frac{e^{-\beta U^{new} + \beta \mu (N+1)} d\mathbf{r}^{N+1}}{\Xi \Lambda^{d(N+1)}} \left[\frac{1}{2} \times \frac{1}{N+1} \times \min(1, \frac{1}{\chi}) \right]$$

Limiting distribution $\pi(\mathbf{r}^N) = \frac{1}{\Xi} \frac{1}{\Lambda^{dN}} e^{-\beta U(\mathbf{r}^N) + \beta \mu N} d\mathbf{r}^N$

Insertion/Deletion Trial Move

3. Analysis of Detailed Balance

Forward-step trial probability $\frac{1}{2} \times \frac{d\mathbf{r}}{V} \times \min(1, \chi)$

Reverse-step trial probability $\frac{1}{2} \times \frac{1}{N+1} \times \min(1, \frac{1}{\chi})$

Detailed balance

$$\frac{\pi_i}{\Xi \Lambda^{dN}} \pi_{ij} = \frac{\pi_j}{\Xi \Lambda^{d(N+1)}} \pi_{ji}$$

$$\frac{e^{-\beta U^{old} + \beta \mu N} d\mathbf{r}^N}{\Xi \Lambda^{dN}} \left[\frac{1}{2} \times \frac{d\mathbf{r}}{V} \times \min(1, \chi) \right] = \frac{e^{-\beta U^{new} + \beta \mu (N+1)} d\mathbf{r}^{N+1}}{\Xi \Lambda^{d(N+1)}} \left[\frac{1}{2} \times \frac{1}{N+1} \times \min(1, \frac{1}{\chi}) \right]$$

$$\frac{1}{V} e^{-\beta U^{old}} \chi = \frac{1}{\Lambda(N+1)} e^{-\beta U^{new} - \beta \mu}$$

Remember

insert: $(N+1) = N_{old} + 1$

delete: $(N+1) = N_{old}$

$$\chi = \frac{V}{\Lambda(N+1)} e^{-\beta(U^{new} - U^{old}) + \beta \mu}$$

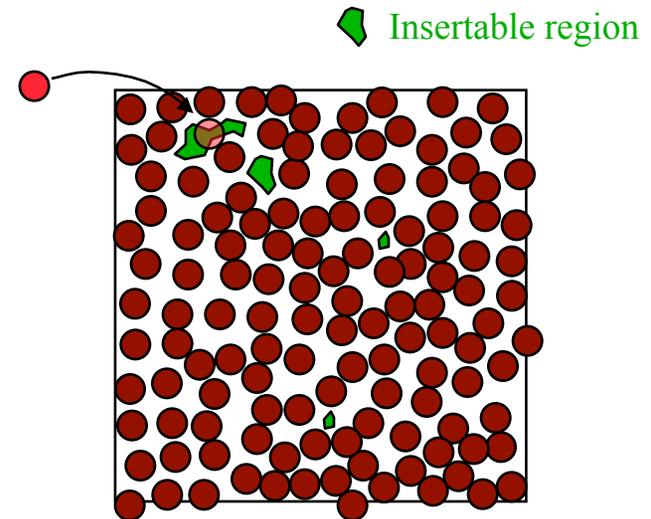
Acceptance probability

Biased Insertion/Deletion Trial Move

1. Specification

○ Trial-move algorithm. Choose with equal probability:

- *Insertion*
 - identify region where insertion will not lead to overlap
 - let the volume of this region be ϵV
 - place randomly somewhere in this region
- *Deletion*
 - select any molecule and delete it



Biased Insertion/Deletion Trial Move

2. Analysis of Trial Probabilities

- Detailed specification of trial moves and and probabilities

Event [reverse event]	Probability [reverse probability]	
Select insertion trial [select deletion trial]	$\frac{1}{2}$ [$\frac{1}{2}$]	<p><i>Forward-step trial probability</i> $\frac{1}{2} \times \frac{dr}{\epsilon V} \times \min(1, \chi)$</p> <p><i>Reverse-step trial probability</i> $\frac{1}{2} \times \frac{1}{N+1} \times \min(1, \frac{1}{\chi})$</p>
Place molecule at \mathbf{r}_{N+1} [delete molecule N+1]	$dr/(\epsilon V)$ [$1/(N+1)$]	
Accept move [accept move]	$\min(1, \chi)$ [$\min(1, 1/\chi)$]	<p>Only difference from unbiased algorithm</p>

Biased Insertion/Deletion Trial Move

3. Analysis of Detailed Balance

Detailed balance

$$\pi_i \pi_{ij} = \pi_j \pi_{ji}$$

$$\frac{e^{-\beta U^{old} + \beta \mu N} d\mathbf{r}^N}{\Xi \Lambda^{dN}} \left[\frac{1}{2} \times \frac{d\mathbf{r}}{\varepsilon V} \times \min(1, \chi) \right] = \frac{e^{-\beta U^{new} + \beta \mu (N+1)} d\mathbf{r}^{N+1}}{\Xi \Lambda^{d(N+1)}} \left[\frac{1}{2} \times \frac{1}{N+1} \times \min(1, \frac{1}{\chi}) \right]$$

Remember

insert: $(N+1) = N_{old} + 1$

delete: $(N+1) = N_{old}$

$$\chi = \frac{\varepsilon V}{\Lambda(N+1)} e^{-\beta(U^{new} - U^{old}) + \beta \mu}$$

Acceptance probability

- ε must be computed even when doing a deletion, since χ depends upon it
 - for deletion, ε is computed for configuration after molecule is removed
 - for insertion, ε is computed for configuration before molecule is inserted

Biased Insertion/Deletion Trial Move

4. Comments

- Advantage is gained when ε is small and $e^{\beta\mu}$ is large
 - *for hard spheres near freezing*
 - $\beta\mu + \ln(V / \Lambda N) : 16$ (difficult to accept deletion without bias)
 - $\varepsilon : 10^{-7}$ (difficult to find acceptable insertion without bias)
 - $\chi : 1$

$$\chi = \varepsilon e^{\beta\mu} \frac{V}{\Lambda(N+1)} e^{-\beta\Delta U}$$

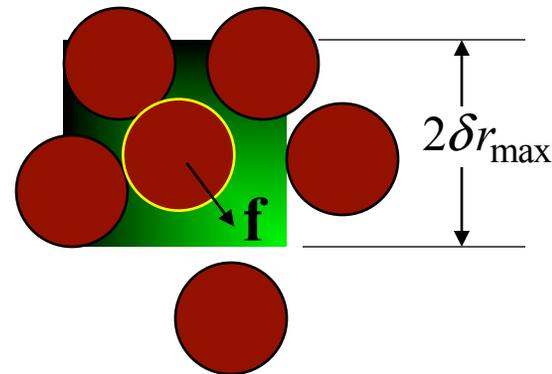
- Identifying and characterizing (computing ε) the non-overlap region may be difficult

Force-Bias Trial Move

1. Specification

○ Move atom in preferentially in direction of lower energy

- select displacement $\delta\mathbf{r}$ in a cubic volume centered on present position
- within this region, select with probability



$$p(\delta\mathbf{r}) = \frac{\exp[+\lambda\beta\mathbf{f} \cdot \delta\mathbf{r}]}{C(\mathbf{f})} = \frac{e^{\lambda\beta f_x \delta r_x} e^{\lambda\beta f_y \delta r_y}}{c_x c_y}$$

Favors δr_y in same direction as f_y

- $C = c_x c_y$ is a normalization constant

$$c_x = \int_{-\delta r_{\max}}^{+\delta r_{\max}} e^{\lambda\beta f_x \delta r_x} d(\delta r_x) = \frac{\sinh(\lambda\beta f_x \delta r_{\max})}{\lambda\beta f_x}$$

An Aside: Sampling from a Distribution

○ Rejection method for sampling from a complex distribution $p(x)$

- write $p(x) = Ca(x)b(x)$
 - $a(x)$ is a simpler distribution
 - $b(x)$ lies between zero and unity
- *recipe*
 - generate a uniform random variate U on $(0,1)$
 - generate a variate X on the distribution $a(x)$
 - if $U < b(X)$ then keep X
 - if not, try again with a new U and X

○ We wish to sample from $p(x) = e^{qx}$ for $x = (-\delta, +\delta)$

- we know how to sample on $e^{q(x-x_0)}$ for $x = (x_0, \infty)$
 - $x = x_0 - q \ln[U(0,1)]$
- use rejection method with
 - $a(x) = e^{q(x-\delta)}$
 - $b(x) = 0$ for $x < -\delta$ or $x > +\delta$; 1 otherwise
 - i.e., sample on $a(x)$ and reject values outside desired range

Force-Bias Trial Move

2. Analysis of Trial Probabilities

- Detailed specification of trial moves and and probabilities

Event [reverse event]	Probability [reverse probability]
Select molecule k [select molecule k]	1/N [1/N]
Move to \mathbf{r}^{new} [move back to \mathbf{r}^{old}]	$p^{\text{old}}(\delta\mathbf{r})$ [$p^{\text{new}}(-\delta\mathbf{r})$]
Accept move [accept move]	$\min(1, \chi)$ [$\min(1, 1/\chi)$]

Forward-step trial probability $\frac{p^{\text{old}}(\delta\mathbf{r})d\mathbf{r}}{N} \times \min(1, \chi)$

Reverse-step trial probability $\frac{p^{\text{new}}(-\delta\mathbf{r})d\mathbf{r}}{N} \times \min(1, \frac{1}{\chi})$

Force-Bias Trial Move

3. Analysis of Detailed Balance

Forward-step trial probability $\frac{p^{old}(\delta\mathbf{r})d\mathbf{r}}{N} \times \min(1, \chi)$

Reverse-step trial probability $\frac{p^{new}(-\delta\mathbf{r})d\mathbf{r}}{N} \times \min(1, \frac{1}{\chi})$

$$p(\delta\mathbf{r}) = \frac{\exp[+\lambda\beta\mathbf{f} \cdot \delta\mathbf{r}]}{C(\mathbf{f})} = \frac{e^{\lambda\beta f_x \delta r_x} e^{\lambda\beta f_y \delta r_y}}{c_x c_y}$$

Detailed balance

$$\frac{e^{-\beta U^{old}} d\mathbf{r}^N}{Z_N} \left[\frac{\pi_i}{N} \times \frac{e^{+\lambda\beta\mathbf{f}^{old} \cdot \delta\mathbf{r}}}{C(\mathbf{f}^{old})} \times \min(1, \chi) \right] = \frac{e^{-\beta U^{new}} d\mathbf{r}^N}{Z_N} \left[\frac{\pi_j}{N} \times \frac{e^{-\lambda\beta\mathbf{f}^{new} \cdot \delta\mathbf{r}}}{C(\mathbf{f}^{new})} \times \min(1, \frac{1}{\chi}) \right]$$

Limiting distribution $\pi(\mathbf{r}^N) d\mathbf{r}^N = \frac{1}{Z_N} e^{-\beta U(\mathbf{r}^N)} d\mathbf{r}^N$

Force-Bias Trial Move

3. Analysis of Detailed Balance

Forward-step trial probability $\frac{p^{old}(\delta\mathbf{r})d\mathbf{r}}{N} \times \min(1, \chi)$

Reverse-step trial probability $\frac{p^{new}(-\delta\mathbf{r})d\mathbf{r}}{N} \times \min(1, \frac{1}{\chi})$

$$p(\delta\mathbf{r}) = \frac{\exp[+\lambda\beta\mathbf{f} \cdot \delta\mathbf{r}]}{C(\mathbf{f})} = \frac{e^{\lambda\beta f_x \delta r_x} e^{\lambda\beta f_y \delta r_y}}{c_x c_y}$$

Detailed balance

$$\frac{\pi_i}{Z_N} \left[\frac{1}{N} \times \frac{e^{+\lambda\beta\mathbf{f}^{old} \cdot \delta\mathbf{r}}}{C(\mathbf{f}^{old})} \times \min(1, \chi) \right] = \frac{\pi_j}{Z_N} \left[\frac{1}{N} \times \frac{e^{-\lambda\beta\mathbf{f}^{new} \cdot \delta\mathbf{r}}}{C(\mathbf{f}^{new})} \times \min(1, \frac{1}{\chi}) \right]$$

$$\frac{1}{C(\mathbf{f}^{old})} e^{-\beta U^{old} + \lambda\beta\mathbf{f}^{old} \cdot \delta\mathbf{r}} \chi = \frac{1}{C(\mathbf{f}^{new})} e^{-\beta U^{new} - \lambda\beta\mathbf{f}^{new} \cdot \delta\mathbf{r}}$$

$$\chi = \frac{C(\mathbf{f}^{old})}{C(\mathbf{f}^{new})} e^{-\beta(U^{new} - U^{old}) - \lambda\beta(\mathbf{f}^{new} + \mathbf{f}^{old}) \cdot \delta\mathbf{r}}$$

Acceptance probability

Force-Bias Trial Move

4. Comments

- Necessary to compute force both before and after move

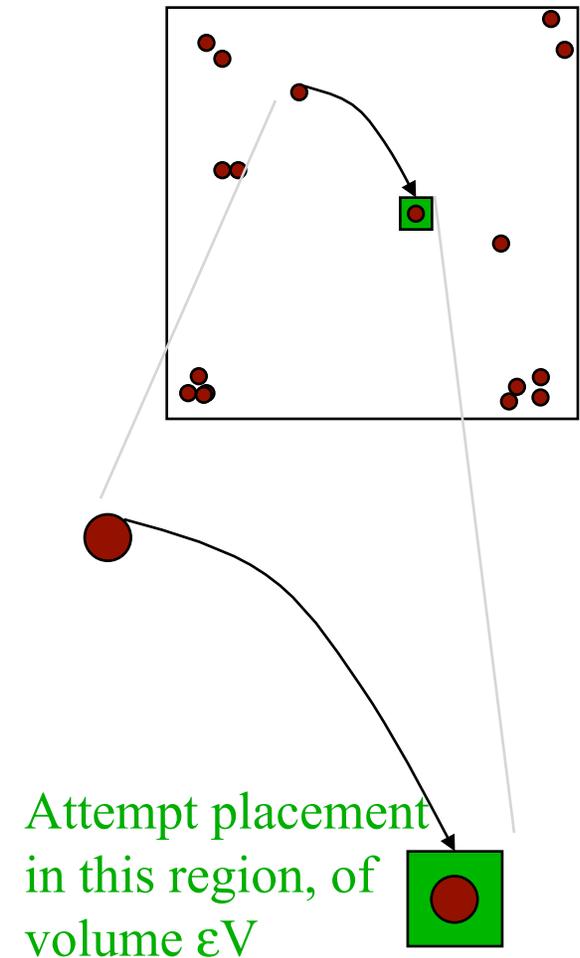
$$\chi = \frac{C(\mathbf{f}^{old})}{C(\mathbf{f}^{new})} e^{-\beta(U^{new} - U^{old}) - \lambda\beta(\mathbf{f}^{new} + \mathbf{f}^{old}) \cdot \delta\mathbf{r}}$$

- From definition of force $\mathbf{f} = -\nabla U$
 - $U^{new} \approx U^{old} - \frac{1}{2}(\mathbf{f}^{new} + \mathbf{f}^{old}) \cdot \delta\mathbf{r}$
 - $\lambda = 1/2$ makes argument of exponent nearly zero
 - $\lambda = 0$ reduces to unbiased case
- Force-bias makes Monte Carlo more like molecular dynamics
 - *example of hybrid MC/MD method*
- Improvement in convergence by factor or 2-3 observed
 - *worth the effort?*

Association-Bias Trial Move

1. Specification

- Low-density, strongly attracting molecules
 - *when together, form strong associations that take long to break*
 - *when apart, are slow to find each other to form associations*
 - *performance of simulation is a problem*
- Perform moves that put one molecule preferentially in vicinity of another
 - *suffer overlaps, maybe 50% of time*
 - *compare to problem of finding associate only 1 time in (say) 1000*
- Must also preferentially attempt reverse move

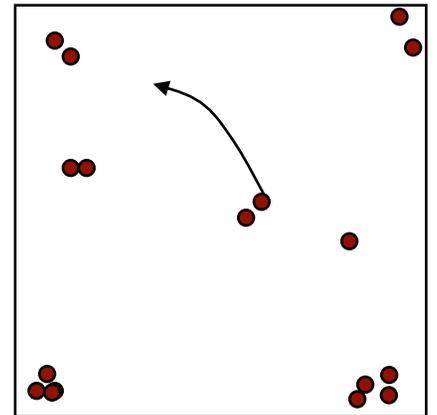
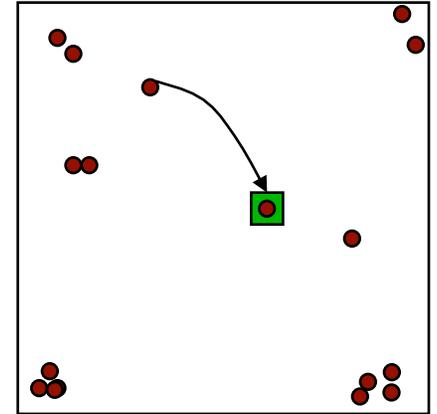


Association-Bias Trial Move

1. Specification

○ With equal probability, choose a move:

- *Association*
 - select a molecule that is not associated
 - select another molecule (associated or not)
 - put first molecule in volume eV in vicinity of second
- *Dis-association*
 - select a molecule that is associated
 - move it to a random position anywhere in the system



Association-Bias Trial Move

2. Analysis of Trial Probabilities

○ Detailed specification of trial moves and and probabilities

Event [reverse event]	Probability [reverse probability]
Select molecule k [select molecule k]	$1/N_{\text{un}}$ [$1/(N_{\text{assoc}}+1)$]
Move to \mathbf{r}^{new} [move back to \mathbf{r}^{old}]	$1/(N_{\text{assoc}}\epsilon V)$ (*) [$1/V$]
Accept move [accept move]	$\min(1, \chi)$ [$\min(1, 1/\chi)$]

Forward-step trial probability $\frac{1}{N_u N_a \epsilon V} \times \min(1, \chi)$

Reverse-step trial probability $\frac{1}{(N_a + 1)V} \times \min(1, \frac{1}{\chi})$

(*) incorrect

Association-Bias Trial Move

3. Analysis of Detailed Balance

Forward-step trial probability $\frac{1}{N_u N_a \epsilon V} \times \min(1, \chi)$

Reverse-step trial probability $\frac{1}{(N_a + 1)V} \times \min(1, \frac{1}{\chi})$

Detailed balance

$$\pi_i \pi_{ij} = \pi_j \pi_{ji}$$

$$\frac{e^{-\beta U^{old}}}{Z_N} \cancel{d\mathbf{r}^N} \left[\frac{1}{N_u N_a \epsilon V} \times \min(1, \chi) \right] = \frac{e^{-\beta U^{new}}}{Z_N} \cancel{d\mathbf{r}^N} \left[\frac{1}{(N_a + 1)V} \times \min(1, \frac{1}{\chi}) \right]$$

$$\chi = \frac{N_a}{(N_a + 1)} N_u \epsilon e^{-\beta(U^{new} - U^{old})}$$

Acceptance probability

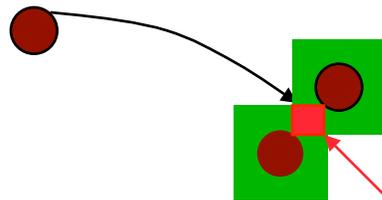
Association-Bias Trial Move

4. Comments

- This is incorrect!

$$\chi = \frac{N_a}{N_a + 1} N_u \epsilon e^{-\beta(U^{new} - U^{old})}$$

- Need to account for full probability of positioning in \mathbf{r}^{new}



This region has extra probability of being selected (in vicinity of two molecules)

- *must look in local environment of trial position to see if it lies also in the neighborhood of other atoms*

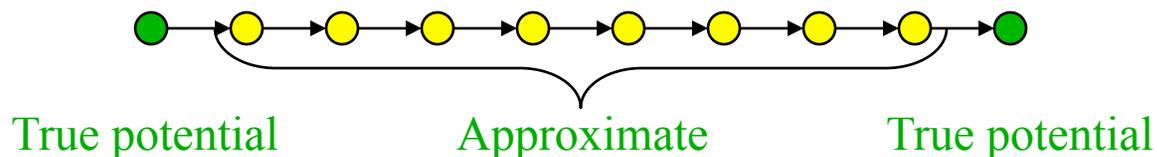
→ add a $1/\epsilon V$ for each atom

- Algorithm requires to identify or keep track of number of associated/unassociated molecules

Using an Approximation Potential

1. Specification

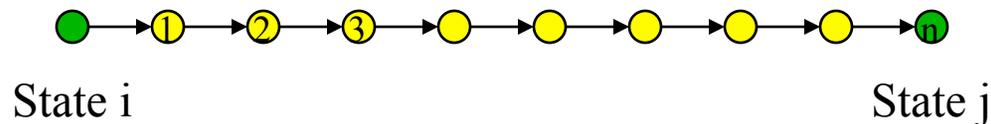
- Evaluating the potential energy is the most time-consuming part of a simulation
- Some potentials are especially time-consuming, e.g.
 - *three-body potentials*
 - *Ewald sum*
- Idea:
 - *move system through Markov chain using an approximation to the real potential (cheaper to compute)*
 - *at intervals, accept or reject entire subchain using correct potential*



Approximation Potential

2. Analysis of Trial Probabilities

- What are π_{ij} and π_{ji} ?



- Given that each elementary Markov step obeys detailed balance for the approximate potential...
 - ...one can show that the “super-step” $i \rightarrow j$ also obeys detailed balance (for the approximate potential)
 - $\pi_i^a \pi_{ij}^{(n)} = \pi_j^a \pi_{ji}^{(n)}$
 - very hard to analyze without this result
 - would have to consider all paths from i to j to get transition probability

Approximation Potential

3. Analysis of Detailed Balance

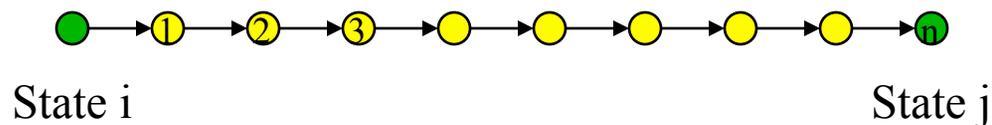
- Formulate acceptance criterion to satisfy detailed balance for the real potential

$$\pi_i \pi_{ij}^{(n)} \min(1, \chi) = \pi_j \pi_{ji}^{(n)} \min(1, 1/\chi)$$

$$\pi_i \left(\frac{\pi_j^a}{\pi_i^a} \pi_{ji}^{(n)} \right) \min(1, \chi) = \pi_j \pi_{ji}^{(n)} \min(1, 1/\chi)$$

Approximate-potential
detailed balance

$$\pi_i^a \pi_{ij}^{(n)} = \pi_j^a \pi_{ji}^{(n)}$$



Approximation Potential

3. Analysis of Detailed Balance

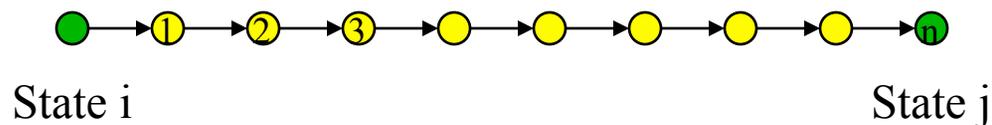
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Approximate-potential
detailed balance

$$\pi_i^a \pi_{ij}^{(n)} = \pi_j^a \pi_{ji}^{(n)}$$



Approximation Potential

3. Analysis of Detailed Balance

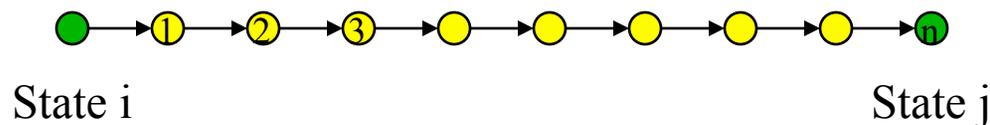
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$$\chi = \frac{\pi_j}{\pi_j^a} \times \frac{\pi_i^a}{\pi_i}$$

Close to 1 if approximate potential is good description of true potential



Summary

- Good Monte Carlo keeps the system moving among a wide variety of states
- At times sampling of wide distribution is not done well
 - *many states of comparable probability not easily reached*
 - *few states of high probability hard to find and then escape*
- Biasing the underlying transition probabilities can remedy problem
 - *add bias to underlying TPM*
 - *remove bias in acceptance step so overall TPM is valid*
- Examples
 - *insertion/deletion bias in GCMC*
 - *force bias*
 - *association bias*
 - *using an approximate potential*